Annales Mathematicae et Informaticae 39 (2012) pp. 173–191

Proceedings of the Conference on Stochastic Models and their Applications Faculty of Informatics, University of Debrecen, Debrecen, Hungary, August 22–24, 2011

# Renewal theorems in the case of attraction to the stable law with characteristic exponent smaller than unity<sup>\*</sup>

#### S. V. Nagaev

Sobolev Institute of Mathematics, Novosibirsk nagaev@math.nsc.ru, nagaevs@hotmail.com, nagaevs@academ.org

Dedicated to Mátyás Arató on his eightieth birthday

## 1. Introduction

Let X be a non-negative integer-valued random variable,  $p_n = \mathbf{P}(X = n)$ . Put  $S_n = \sum_{j=1}^n X_j$ ,  $n \ge 1$ , where  $X_j$  are i.i.d. random variables which have the same distribution as X. In what follows we assume that  $S_0 = 0$ . Let  $u_n = \sum_{k=0}^{\infty} \mathbf{P}(S_k = n)$  be the renewal probability at the instant n. Put  $f(z) = \sum_{k=0}^{\infty} p_k z^k$ . If g(z) is an analytical function in some neighbourhood of zero, we denote the coefficient at  $z^n$  in Taylor series for g(z) by  $C_n(g(z))$ .

In 1963 Garsia and Lamperti [1] proved that under the condition

$$\mathbf{P}(X > n) \sim L(n)n^{-\alpha},\tag{1.1}$$

where L(x) is a slowly-varying function, the asymptotic formula

$$u_n \sim \frac{\sin \pi \alpha}{\pi} L^{-1}(n) n^{\alpha - 1}, \qquad (1.2)$$

is valid, provided  $1/2 < \alpha < 1$ . The relation  $a_n \sim b_n$  here and below indicates that  $\lim_{n \to \infty} a_n/b_n = 1$ .

In 1968 Williamson [3] extended Garsia-Lamperti's result to the case that X belongs to the domain of attraction of a non-degenerate d - dimensional stable law with characteristic exponent  $\alpha$ ,  $d/2 < \alpha < \min(d; 2)$ .

To prove (1.2) Garsia and Lamperti used the purely analytical method based on analysis of behavior of the generating function f(z) on the unit circle. On the

<sup>\*</sup>This work was supported by RFBR 09-01-00048-a.

contrary, Williamson's approach is probabilistic with the local limit theorem by Rvacheva [4] as the starting point.

As to case  $0 < \alpha \leq 1/2$ , formula (1.2), generally speaking, is not true if we restrict our selves to condition (1.1). Corresponding counter-example is given in [3]. The point is that in the case  $0 < \alpha \leq 1/2$  the existence of lacunas in the sequence  $p_n$  influences on the behavior of  $u_n$ . Therefore, additional constraints are necessary to provide the validity of (1.2). One such constraint was suggested by De Bruijn and Erdos [2] before [1] appeared, namely,

$$p_{n-1}p_{n+1} > p_n^2, (1.3)$$

i.e. the sequence  $\ln p_n$  is convex. Williamson [3] noticed that (1.2) remains true if the sequence  $p_n$  does not increase beginning with some number n. This condition is weaker than (1.3).

In the present work we use the condition

$$p_n \sim \frac{l(n)}{n^{1+\alpha}}, \ 0 < \alpha < 1, \tag{1.4}$$

where the function l(x) is slowly varying. Notice that condition (1.1) with  $L(n) = \alpha^{-1}l(n)$  follows from (1.4) (see Lemma 2.1 below). Condition (1.4) is discussed in our previous paper [5], namely, it is shown therein that if above-mentioned Williamson's condition is fulfilled, then (1.4) hold.

**Theorem 1.1.** If condition (1.4) holds, then

$$u_n \sim c(\alpha) \frac{\mathbf{P}(X=n)}{\mathbf{P}^2(X \ge n)} \sim \frac{\alpha^2 c(\alpha)}{l(n)n^{1-\alpha}},\tag{1.5}$$

where  $c(\alpha) = \sin \pi \alpha / \pi \alpha$ .

The extreme case  $p_n \sim n^{-1}l(n)$  is studied in [5]. It turns out that under this condition  $u_n \sim \mathbf{P}(X=n)/\mathbf{P}^2(X \ge n)$ . Since  $c(\alpha) \to 1$  as  $\alpha \to 0$ , it implies that representation

$$u_n \sim c(\alpha) \frac{\mathbf{P}(X=n)}{\mathbf{P}^2(X \ge n)},$$

which is given in Theorem 1.1 is stable as  $\alpha \to 0$ . However, we can not say this about the relation  $u_n \sim \alpha^2 c(\alpha)/l(n)n^{1-\alpha}$ .

In proving Theorem 1.1 we apply the same approach as in [5]. However, to realize it was found more difficult in this case.

*Remark.* In [6] the renewal theorem is proved under condition that (1.1) holds and

$$p_n < c\mathbf{P}(X > n)n^{-1}$$

using Williamson's method. The proof is based on the following statement: Assume that F(0) = 0 and (2.1) holds. Then for all  $n \ge 1, z$  large enough and  $x \ge z$ 

$$\mathbf{P}\{S_n \ge x, M_n \le z\} \le \{cz/x\}^{x/z},$$

where  $M_n = \max\{X_1, X_2, ..., X_n\}$  and  $S_n = \sum_{i=1}^{n} X_i$  (see Lemma 2 in [6]).

The author of [6] asserts that this lemma is an immediate consequence of the inequality

$$\mathbf{P}(S_n \ge x) \le \sum_{i=1}^{n} \mathbf{P}(X_i > y_i) + (eA_t^+ / xy^{t-1})^{x/y},$$

where  $S_n = \sum_{j=1}^n X_j$ ,  $X_j$  are independent random variables,  $y > \max_i y_i$ ,  $A_t^+ = \sum_{j=1}^n \{X_j^t; X_j > 0\}$ , 0 < t < 1 (see Corollary 1.5 in [7]). If  $X_j$  are i.i.d. equal to X by distribution, then

$$\mathbf{P}(S_n \ge x) \le n\mathbf{P}(X > y) + \left(\frac{en\mathbf{E}\{X^t; X > 0\}}{xy^{t-1}}\right)^{x/y}$$

If  $X \leq y$ , then

$$\mathbf{E}\{X^t; X > 0\} \le y^t.$$

Consequently, in this case

$$\mathbf{P}(S_n \ge x) \le \left(\frac{eny}{x}\right)^{x/y}.$$

This inequality differs from the inequality stated in [6] by the presence of n in the right-hand side. Thus, Lemma 2 of [6] does not follow from Corollary 1.5 of [7]), and, therefore, the former can not be considered as being proved.

Let 
$$h_n = \sum_{k=0}^{\infty} n^{-1} \mathbf{P}(S_k = n).$$

**Theorem 1.2.** If condition (1.4) holds, then

$$h_n \sim \frac{\alpha}{n}.\tag{1.6}$$

Notice that  $h_n$  is the derivative of the measure  $\nu(A) := \sum_{k \in A} h_k$  with respect to the counting measure. The measure  $\nu(A)$  is a particular case of so called harmonic renewal measure. Recall that the measure  $\nu(\cdot) = \sum_{1}^{\infty} n^{-1}F_n(\cdot)$ , where  $F_n$  is *n*-th convolution of any distribution F on  $\mathbf{R}^+$  is said to be harmonic renewal measure associated with F. In our case the distribution F is concentrated on the lattice of non-negative integers. The harmonic renewal function is defined by the equality  $H(x) = \nu([0, x))$ .

The next statement concerning the asymptotic behavior of H(n) as  $n \to \infty$  follows from Theorem 1.2.

**Corollary 1.3.** If condition (1.4) holds, then

$$H(n) \sim \alpha \ln n. \tag{1.7}$$

The asymptotic behavior of H(x) for  $x \to \infty$  is studied in [9, 10, 11, 12]. The case that F attracts to a stable law is considered in [9], namely, it is proved therein that under the condition  $1 - F(x) \sim x^{-\alpha}L(x)$ 

$$\lim_{x \to \infty} (H(x) - \alpha \ln x + \ln L(x)) = \alpha \mathbf{C} - \ln \Gamma(1 - \alpha),$$

where **C** is the Euler constant,  $\Gamma(\cdot)$  is the gamma function. Of course, the last assertion is sharper than (1.7). Formula (1.7) is presented by reason of simplicity of proving.

#### 2. Auxiliary results

**Lemma 2.1.** For any  $\alpha > 0$ 

$$\sum_{k=n}^{\infty} \frac{l(k)}{k^{\alpha+1}} \sim \int_{n}^{\infty} \frac{l(y)}{y^{\alpha+1}} dy.$$
(2.1)

*Proof.* Put  $p(x) = l(x)/x^{\alpha+1}$ . Obviously,

$$\inf_{n \le y \le n+1} \frac{p(y)}{p(n)} \le \frac{1}{p(n)} \int_{n}^{n+1} p(y) dy \le \sup_{n \le y \le n+1} \frac{p(y)}{p(n)}.$$
(2.2)

It is easily seen that for every  $n \le y \le n+1$ 

$$\left(\frac{n}{n+1}\right)^{\alpha+1} \inf_{\substack{n \le y \le n+1}} \frac{l(y)}{l(n)} \le \frac{p(y)}{p(n)} \le \sup_{\substack{n \le y \le n+1}} \frac{l(y)}{l(n)}.$$
(2.3)

In what follows we need Kamarata's representation

$$l(x) = a(x) \exp\left\{\int_{1}^{x} \frac{\epsilon(u)}{u} du\right\}, \ x \ge 1,$$
(2.4)

where  $\lim_{n \to \infty} \varepsilon(u) = 0$ ,  $\lim_{x \to \infty} a(x) = a$ ,  $0 < a < \infty$ . Hence,

$$\frac{l(y)}{l(n)} = \frac{a(y)}{a(n)} \exp\bigg\{\int_{n}^{y} \frac{\epsilon(u)}{u} du\bigg\}.$$

Obviously,

$$\lim_{n \to \infty} \sup_{n \le y \le n+1} \left| \int_{n}^{y} \frac{\varepsilon(u)}{u} du \right| = 0.$$

It follows from last two relations that

$$\lim_{n \to \infty} \sup_{n \le y \le n+1} \left| \frac{l(y)}{l(n)} - 1 \right| = 0.$$
 (2.5)

Combining (2.2), (2.3) and (2.5), we have

$$\lim_{n \to \infty} \frac{1}{p(n)} \int_{n}^{n+1} p(y) dy = 1.$$
 (2.6)

It is easily seen that

$$\inf_{k \ge n} \frac{1}{p(k)} \int_{k}^{k+1} p(y) dy \le \frac{\int_{n}^{\infty} p(y) dy}{\sum_{k=n}^{\infty} p(k)} \le \sup_{k \ge n} \frac{1}{p(k)} \int_{k}^{k+1} p(y) dy.$$
(2.7)

The conclusion of the Lemma follows from (2.6) and (2.7).

**Lemma 2.2.** For any  $\alpha > 0$ 

$$\int_{x}^{\infty} \frac{l(y)}{y^{\alpha+1}} dy \sim \frac{l(x)}{\alpha x^{\alpha}}.$$
(2.8)

*Proof.* By using (2.4), we have

$$\int_{x}^{\infty} \frac{l(y)}{y^{\alpha+1}} dy \sim \int_{x}^{\infty} \frac{l_0(y)}{y^{\alpha+1}} dy,$$
(2.9)

where

$$l_0(y) = \exp\bigg\{\int_1^y \frac{\varepsilon(u)}{u} du\bigg\}.$$
(2.10)

Integrating by parts, we conclude that

$$\int_{x}^{\infty} \frac{l_0(y)}{y^{\alpha+1}} dy = \frac{l_0(x)}{\alpha x^{\alpha}} + \frac{1}{\alpha} \int_{x}^{\infty} \frac{\varepsilon(u)l_0(y)}{y^{\alpha+1}} dy$$

$$= \frac{l_0(x)}{\alpha x^{\alpha}} (1+o(1)) = \frac{l(x)}{\alpha x^{\alpha}} (1+o(1)).$$
(2.11)

The desired result follows from (2.9) and (2.11).

Note that (2.8) can be deduced from the asymptotic formula

$$\int_{\alpha}^{\infty} f(t)l(xt)dt \sim l(x) \int_{\alpha}^{\infty} f(t)dt,$$

where  $\alpha > 0$ , and  $f(t)t^{\eta}$ ,  $\eta > 0$ , is integrable (see [8], Theorem 2.6), but not immediately. For this purpose one needs to make the change of variables y = xtin the integral  $\int_{x}^{\infty} y^{-\alpha-1}l(y)dy$ . On the other hand, the method which is used in proving Lemma 2.2 allows to obtain very easily the statement the above mentioned Theorem 2.6 of [8].

**Corollary 2.3.** Under condition (1.4)

$$\mathbf{P}(X \ge n) \sim \frac{l(n)}{\alpha n^{\alpha}}.$$
(2.12)

Proof. Evidently,

$$\inf_{k \ge n} \frac{l(k)}{k^{\alpha+1} p_k} \le \frac{\sum_{k=n}^{\infty} l(k) k^{-\alpha-1}}{\sum_{k=n}^{\infty} p_k} \le \sup_{k \ge n} \frac{l(k)}{k^{\alpha+1} p_k}.$$

Hence, by (2.7)

$$\mathbf{P}(X \ge n) = \sum_{k \ge n} p_k \sim \sum_{k \ge n} \frac{l(k)}{k^{\alpha+1}} \sim \frac{l(n)}{\alpha n^{\alpha}}$$

which was to be proved.

**Lemma 2.4.** For any  $\alpha < 1$ 

$$\sum_{k=1}^{n} \frac{l(k)}{k^{\alpha}} \sim \frac{l(n)}{1-\alpha} n^{1-\alpha}.$$
(2.13)

*Proof.* Let  $l_0(x)$  be defined by (2.10). Since  $l_0(x) \sim l(x)$ ,

$$\sum_{k=1}^{n} \frac{l_0(k)}{k^{\alpha}} \sim \sum_{k=1}^{n} \frac{l(k)}{k^{\alpha}}.$$
(2.14)

Indeed,

$$1 - \varepsilon < \frac{\sum\limits_{k=n(\varepsilon)}^{n} k^{-\alpha} l_0(k)}{\sum\limits_{k=n(\varepsilon)}^{n} k^{-\alpha} l(k)} < 1 + \varepsilon$$

if  $n(\varepsilon)$  is such that for  $x > n(\varepsilon)$ 

$$1 - \varepsilon < \frac{l_0(x)}{l(x)} < 1 + \varepsilon.$$

It is easily seen that

$$\lim_{n \to \infty} \sum_{k=n(\epsilon)}^{n} k^{-\alpha} l(k) = \infty.$$

Therefore for sufficiently large n

$$1 - 2\varepsilon < \frac{\sum\limits_{k=n(\varepsilon)}^{n} k^{-\alpha} l_0(k)}{\sum\limits_{k=n(\varepsilon)}^{n} k^{-\alpha} l(k)} < 1 + 2\varepsilon.$$

Since  $\varepsilon$  can be made as small as we wish, hence the validity of (2.14) follows. By applying the Abel transform, we get

$$\sum_{k=1}^{n} \frac{l_0(k)}{k^{\alpha}} = l_0(n) \sum_{k=1}^{n} k^{-\alpha} + \sum_{k=1}^{n-1} (l_0(k) - l_0(k+1)) \sum_{j=1}^{k} j^{-\alpha}.$$
 (2.15)

It is easily seen that

$$l_0(k) - {}_0(k+1) = l_0(k) \left( 1 - \exp\left\{ \int_k^{k+1} \frac{\varepsilon(u)}{u} du \right\} \right).$$

Hence

$$|l_0(k) - {}_0(k+1)| < l_0(k) \Big| \int_k^{k+1} \frac{\varepsilon(u)}{u} du \Big| = o(l_0(k)k^{-1}).$$
(2.16)

Further,

$$\sum_{k=1}^{n} k^{-\alpha} \sim \frac{n^{1-\alpha}}{1-\alpha}.$$
(2.17)

It follows from (2.16) and (2.17)

$$\sum_{k=1}^{n-1} (l_0(k) - l_0(k+1)) \sum_{j=1}^k j^{-\alpha} = o\left(\sum_{k=1}^n \frac{l_0(k)}{k^{\alpha}}\right).$$
(2.18)

Combining (2.15)–(2.17), we conclude that

$$\sum_{k=1}^{n} l_0(k) k^{-\alpha} \sim \frac{l_0(n)}{1-\alpha} n^{1-\alpha}.$$
(2.19)

From (2.14) and (2.19) the result follows.

Corollary 2.5. Under conditions of Theorem 1.1

$$\sum_{k=1}^{n} \mathbf{P}(X \ge k) \sim \frac{l(n)}{\alpha(1-\alpha)} n^{1-\alpha}.$$
(2.20)

*Proof.* According to Corollary 2.3 for any  $\varepsilon > 0$  there exists  $n(\varepsilon)$  such that for  $n > n(\varepsilon)$ 

$$1 - \varepsilon < \mathbf{P}(X \ge n) / \frac{l(n)}{\alpha n^{\alpha}} < 1 + \varepsilon.$$

Hence,

$$1 - \varepsilon < \sum_{n(\varepsilon) < k \le n}^{n} \mathbf{P}(X \ge k) / \alpha^{-1} \sum_{n(\varepsilon) < k \le n}^{n} \frac{l(k)}{k^{\alpha}} < 1 + \varepsilon.$$

On the other hand, since

$$\lim_{n \to \infty} \sum_{n(\varepsilon) < k \le n} \frac{l(k)}{k^{\alpha}} = \infty$$

for every  $\varepsilon > 0$ ,

$$\sum_{n(\varepsilon) < k \leq n} \frac{l(k)}{k^{\alpha}} \sim \sum_{k=1}^n \frac{l(k)}{k^{\alpha}}, \ \sum_{n(\varepsilon) < k \leq n}^n \mathbf{P}(X \geq k) \sim \sum_{k=1}^n \mathbf{P}(X \geq k).$$

Therefore, for sufficiently large n

$$1 - 2\varepsilon < \alpha \sum_{k=1}^{n} \mathbf{P}(X \ge k) / \sum_{k=1}^{n} \frac{l(k)}{k^{\alpha}} < 1 + 2\varepsilon.$$

Hence, since  $\varepsilon$  is arbitrary, it follows that

$$\sum_{k=1}^{n} \mathbf{P}(X \ge k) \sim \alpha^{-1} \sum_{k=1}^{n} \frac{l(k)}{k^{\alpha}}.$$

To complete the proof it remains to apply Lemma 2.4.

Lemma 2.6. Under conditions of Theorem 1.1

$$1 - f(z) \sim (1 - z)^{\alpha} L\left(\frac{1}{1 - z}\right),$$
 (2.21)

where

$$L(x) = \frac{\Gamma(1-\alpha)}{\alpha} l(x).$$

Proof. First of all,

$$\sum_{k=0}^{n} \mathbf{P}(X > k) z^{k} = \frac{1 - f(z)}{1 - z}.$$

It is easily seen that

$$\mathbf{P}(X > k) \sim \mathbf{P}(X \ge k).$$

Now, using Corollary 2.5 and the Abelian theorem (see, e.g. [13], Ch. XIII, section 5, Th. 5), we have

$$\frac{1-f(z)}{1-z} \sim \frac{\Gamma(2-\alpha)}{\alpha(1-\alpha)} (1-z)^{\alpha-1} L(1-z)$$
  
=  $\alpha^{-1} \Gamma(1-\alpha) (1-z)^{\alpha-1} l\left(\frac{1}{1-z}\right) = (1-z)^{\alpha-1} L\left(\frac{1}{1-z}\right),$ 

which is equivalent to the assertion of the Lemma.

Lemma 2.7. Under conditions of Theorem 1.1

$$\sum_{k=0}^{n} u_k \sim \frac{n^{\alpha}}{\Gamma(\alpha+1)L(n)},\tag{2.22}$$

where L(x) is defined in Lemma 2.6.

Proof. Obviously,

$$u_k = C_k \left(\frac{1}{1 - f(z)}\right).$$

Applying Lemma 2.6 and the Tauberian theorem (see ref. in the proof of Lemma 2.6), we obtain the desired result.  $\Box$ 

The next assertion is borrowed from [5].

Lemma 2.8. The identity

$$nu_n = \sum_{k=0}^{n-1} (n-k)p_{n-k}u_k^{(2)}$$
(2.23)

holds, where  $u_n = \sum_{k=0}^{\infty} \mathbf{P}(S_k = n), \ u_n^{(2)} = \sum_{k=0}^n u_{n-k} u_k.$ 

**Lemma 2.9.** Under condition of Theorem 1.1 there exists the sequence  $\theta_n$  such that  $\lim_{n\to\infty} \theta_n = 1$  and

$$u_n^{(2)} \le \frac{2^{1-\alpha}\theta_n n^{\alpha}}{\Gamma(\alpha+1)L(n)} \max_{n/2 \le k \le n} u_k.$$

$$(2.24)$$

*Proof.* It is easily seen that

$$u_n^{(2)} \le 2 \sum_{0 \le k \le n/2} u_k u_{n-k} \le 2 \max_{n/2 \le k \le n} u_k \sum_{0 \le k \le n/2} u_k$$

To complete the proof it is sufficient to apply Lemma 2.7.

Lemma 2.10. Under conditions of Theorem 1.1

$$\sum_{k=1}^{n} u_k^{(2)} \sim \frac{n^{2\alpha}}{\Gamma(2\alpha+1)L^2(n)},$$
(2.25)

where L(x) is defined in Lemma 2.6.

*Proof.* It is easily seen that

$$u_k^{(2)} = C_k \left(\frac{1}{(1-f(z))^2}\right).$$

According to Lemma 2.6

$$(1-f(z))^{-2} \sim (1-z)^{-2\alpha} L^{-2} \left(\frac{1}{1-z}\right).$$

Applying the Tauberian theorem (see ref. in the proof of Lemma 2.6), we get the desired result.  $\hfill \Box$ 

**Lemma 2.11.** Under conditions of Theorem 1.1 for every fixed 0 < a < b < 1

$$\sum_{na \le k \le nb} l^{-2}(k)k^{2\alpha - 1}(n-k)^{-\alpha} \sim l^{-2}(n)n^{\alpha} \int_{a}^{b} u^{2\alpha - 1}(1-u)^{-\alpha} du.$$
(2.26)

*Proof.* First of all, notice that

$$\ln \frac{l_0(n)}{l_0(k)} = \int_k^n \frac{\varepsilon(u)}{u} du.$$
(2.27)

Consequently,

$$\lim_{n \to \infty} \sup_{na \le k \le nb} \left| \frac{l_0(n)}{l_0(k)} - 1 \right| = 0.$$
(2.28)

This implies that

$$\sum_{na \le k \le nb} l_0^{-2}(k) k^{2\alpha - 1} (n - k)^{-\alpha} \sim l_0^{-2}(n) \sum_{na \le k \le nb} k^{2\alpha - 1} (n - k)^{-\alpha}$$

Hence it follows that

$$\sum_{na \le k \le nb} l^{-2}(k)k^{2\alpha-1}(n-k)^{-\alpha} \sim l^{-2}(n) \sum_{na \le k \le nb} k^{2\alpha-1}(n-k)^{-\alpha}.$$

Further,

$$\sum_{na \le k \le nb} k^{2\alpha - 1} (n - k)^{-\alpha} = n^{\alpha - 1} \sum_{na \le k \le nb} \left(\frac{k}{n}\right)^{2\alpha - 1} \left(1 - \frac{k}{n}\right)^{-\alpha}$$
$$\sim n^{\alpha} \int_{a}^{b} u^{2\alpha - 1} (1 - u)^{-\alpha} du.$$

The result follows from last two relations.

### 3. The proof of Theorem 1.1

Let us write down formula (2.23) in the form

$$nu_{n} = \left(\sum_{0 \le k < \sqrt{n}} + \sum_{\sqrt{n} \le k \le (1-\eta)n} + \sum_{(1-\eta)n < k \le n}\right) (n-k) p_{n-k} u_{k}^{(2)}$$
  
$$\equiv \sum_{1} + \sum_{2} + \sum_{3},$$
(3.1)

where  $0 < \eta < 1$ . For any  $\varepsilon > 0$ , sufficiently large n, and  $k < \sqrt{n}$ 

$$p_{n-k} < (1+\varepsilon)\frac{l(n-k)}{(n-\sqrt{n})^{\alpha+1}}.$$
(3.2)

If  $n - \sqrt{n} \le k \le n$ , then

$$\frac{l_0(n)}{l_o(k)} = \exp\left\{\int_k^n \frac{\varepsilon(u)}{u} du\right\} = 1 + o(\ln n - \ln(n - \sqrt{n})) = 1 + o\left(\frac{1}{\sqrt{n}}\right).$$

Consequently,

$$\max_{n-\sqrt{n}\leq k\leq n} l_0(k) \sim l_0(n).$$
(3.3)

It follows from (3.2) and (3.3) that

$$\sum_{1} = O\left(\frac{l(n)}{n^{\alpha}} \sum_{k=1}^{\lceil \sqrt{n} \rceil} u_{k}^{(2)}\right).$$

By Lemma 2.10

$$\sum_{k=1}^{\left[\sqrt{n}\right]} u_k^{(2)} = O\left(\frac{n^{\alpha}}{l^2(\sqrt{n})}\right).$$
(3.4)

Thus,

$$\sum_{1} = O\left(\frac{1}{l(\sqrt{n})}\right). \tag{3.5}$$

Let us turn to estimating  $\sum_2$ . It is easily seen that

$$\sum_{2} \sim \sum_{\sqrt{n} \le k \le (1-\eta)n} u_k^{(2)} \frac{l_0(n-k)}{(n-k)^{\alpha}} \equiv \sum_{4}.$$
(3.6)

Applying Abel's transformation, we have

$$\sum_{4} \sim \frac{l_0(n - \sqrt{n})^{\alpha}}{(n - \sqrt{n})^{\alpha}} \sum_{\sqrt{n} \le k \le (1 - \eta)n} u_k^{(2)} - \sum_{\sqrt{n} \le k \le (1 - \eta)n} \left( \frac{l_0(n - k - 1)}{(n - k - 1)^{\alpha}} - \frac{l_0(n - k)}{(n - k)^{\alpha}} \right) \sum_{j = [\sqrt{n}]}^k u_j^{(2)}.$$
(3.7)

By Lemma 2.10

$$\sum_{\sqrt{n} \le k \le (1-\eta)n} u_k^{(2)} = \sum_{k \le (1-\eta)n} u_k^{(2)} - \sum_{k < \sqrt{n}} u_k^{(2)} \sim \frac{(1-\eta)^{2\alpha} n^{2\alpha}}{\Gamma(2\alpha+1)L^2(n)}.$$
 (3.8)

Further,

$$\frac{l_0(k)}{k^{\alpha}} - \frac{l_0(k+1)}{(k+1)^{\alpha}} = l_0(k) \left(\frac{1}{k^{\alpha}} - \frac{1}{(k+1)^{\alpha}}\right) + \frac{l_0(k) - l_0(k+1)}{(k+1)^{\alpha}}.$$
(3.9)

Obviously,

$$\frac{1}{k^{\alpha}} - \frac{1}{(k+1)^{\alpha}} \sim \frac{\alpha}{k^{\alpha+1}}.$$
 (3.10)

On the other hand,

$$l_{0}(k+1) - l_{0}(k) = l_{0}(k) \left( \frac{l_{0}(k+1)}{l_{0}(k)} - 1 \right)$$
  
=  $l_{0}(k) \left( \exp\left\{ \int_{k}^{k+1} \frac{\varepsilon(u)}{u} du \right\} - 1 \right) = o\left(\frac{l_{0}(k)}{k}\right).$  (3.11)

It follows from (3.9)–(3.11) that

$$\frac{l_0(k)}{k^{\alpha}} - \frac{l_0(k+1)}{(k+1)^{\alpha}} \sim \frac{\alpha l_0(k)}{k^{\alpha+1}}.$$
(3.12)

Combining (3.6)–(3.8) and (3.12), we obtain

$$\begin{split} \sum_{2} &\sim \frac{(1-\eta)^{2\alpha} l_{0}(n)n^{\alpha}}{\Gamma(2\alpha+1)L^{2}(n)} - \alpha \sum_{\sqrt{n} \leq k \leq (1-\eta)n} \frac{l_{0}(n-k)}{(n-k)^{\alpha+1}} \sum_{j=[\sqrt{n}]}^{k} u_{j}^{(2)} \\ &= \frac{(1-\eta)^{2\alpha} \alpha n^{\alpha}}{\Gamma(1-\alpha)\Gamma(2\alpha+1)a(n)L(n)} \\ &- \alpha \sum_{\sqrt{n} \leq k \leq (1-\eta)n} \frac{l_{0}(n-k)}{(n-k)^{\alpha+1}} \sum_{j=0}^{k} u_{j}^{(2)} + \alpha \sum_{j=0}^{\sqrt{n}-1} u_{j}^{(2)} \sum_{\sqrt{n} \leq k \leq (1-\eta)n} \frac{l_{0}(n-k)}{(n-k)^{\alpha+1}} \\ &= \frac{(1-\eta)^{2\alpha} \alpha n^{\alpha}}{\Gamma(1-\alpha)\Gamma(2\alpha+1)a(n)L(n)} - \alpha \sum_{5} + \alpha \sum_{6}. \end{split}$$
(3.13)

Here  $a(\cdot)$  is a factor in Karamata's representation (2.4) for l(x). In view of (3.4)

$$\sum_{6} = O\left(\frac{l_0(n)}{l_0^2(\sqrt{n})}\right). \tag{3.14}$$

We now proceed to estimating  $\sum_5$ . By Lemma 2.10

$$\sum_{5} \sim c(\alpha) \sum_{\sqrt{n} \le k \le (1-\eta)n} L^{-2}(k) k^{2\alpha} \frac{l_0(n-k)}{(n-k)^{\alpha+1}} \equiv c(\alpha) \sum_{7},$$
(3.15)

where  $c(\alpha) = 1/\Gamma(2\alpha + 1)$ . Applying the Abel transformation, we have

$$\sum_{7} \sim L^{-2}(n)(1-\eta)^{2\alpha} n^{2\alpha} \sum_{\sqrt{n} \le k \le (1-\eta)n} \frac{l_0(n-k)}{(n-k)^{\alpha+1}} - \sum_{\sqrt{n} \le k \le (1-\eta)n} (L^{-2}(k+1)(k+1)^{2\alpha} - L^{-2}(k)k^{2\alpha}) \sum_{j=[\sqrt{n}]}^k \frac{l_0(n-j)}{(n-j)^{\alpha+1}}.$$
 (3.16)

In the same way as (3.12) we deduce that

$$L^{-2}(k+1)(k+1)^{2\alpha} - L^{-2}(k)k^{2\alpha} \sim 2\alpha L^{-2}(k)k^{2\alpha-1}.$$

Hence, denoting the second summand in (3.16) by  $\sum_{8}$ , we obtain

$$\sum_{k} \sim 2\alpha \sum_{\sqrt{n} \le k \le (1-\eta)n} L^{-2}(k) k^{2\alpha-1} \sum_{j=[\sqrt{n}]}^{k} \frac{l_0(n-j)}{(n-j)^{\alpha+1}}$$

$$\sim 2\alpha l_0(n) \sum_{\sqrt{n} \le k \le (1-\eta)n} L^{-2}(k) k^{2\alpha-1} \sum_{j=[\sqrt{n}]}^k (n-j)^{-\alpha-1}.$$
 (3.17)

It is not difficult to check that for  $\sqrt{n} \le k \le (1 - \eta)n$ 

$$\alpha \sum_{j=|\sqrt{n}|} (n-j)^{-\alpha-1} = (n-k)^{-\alpha} - n^{-\alpha} + o(n^{-\alpha}).$$

Consequently,

$$\sum_{k} +2n^{-\alpha} \sum_{\sqrt{n} \le k \le (1-\eta)n} L^{-2}(k) k^{2\alpha-1} \sim 2l_0(n) \sum_{\sqrt{n} \le k \le (1-\eta)n} L^{-2}(k) k^{2\alpha-1} (n-k)^{-\alpha}.$$
(3.18)

We need the identity

$$\sum_{\sqrt{n} \le k \le (1-\eta)n} = \left(\sum_{\sqrt{n} \le k < \varepsilon n} + \sum_{\varepsilon n \le k \le (1-\eta)n}\right) L^{-2}(k) k^{2\alpha - 1} (n-k)^{-\alpha}$$
  
$$\equiv \sum_{9} + \sum_{10}.$$
(3.19)

It is easily seen that

$$\sum_{9} < (1-\varepsilon)^{-\alpha} n^{-\alpha} \sum_{\sqrt{n} \le k \le \varepsilon n} L^{-2}(k) k^{2\alpha-1}.$$

By using Lemma 2.4, we obtain that

$$\sum_{\sqrt{n} \le k \le \varepsilon n} L^{-2}(k) k^{2\alpha - 1} \sim \frac{(\varepsilon n)^{2\alpha}}{2\alpha L^2(n)}.$$

Therefore, for sufficiently large n

$$\sum_{9} < (1-\varepsilon)^{-\alpha} \frac{\varepsilon^{2\alpha} n^{\alpha}}{2\alpha L^{2}(n)}.$$
(3.20)

On the other hand, by Lemma 2.11

$$\sum_{10} \sim L^{-2}(n) n^{\alpha} \int_{\varepsilon}^{1-\eta} u^{2\alpha-1} (1-u)^{-\alpha} du.$$
 (3.21)

It follows from (3.18) - (3.21) that

$$\sum_{8} + \frac{(1-\eta)^{2\alpha} n^{\alpha}}{\alpha L^{2}(n)} \sim \frac{2\alpha^{2} n^{\alpha}}{\Gamma^{2}(1-\alpha)l(n)} \int_{0}^{1-\eta} u^{2\alpha-1} (1-u)^{-\alpha} du.$$
(3.22)

Combining (3.15), (3.16), (3.18) and (3.22) we obtain

$$\sum_{5} \sim \frac{(1-\eta)^{2\alpha} \alpha n^{\alpha}}{\Gamma(1-\alpha)\Gamma(2\alpha+1)L(n)} - \frac{2\alpha^{2}n^{\alpha}}{\Gamma^{2}(1-\alpha)\Gamma(2\alpha+1)l(n)}I(\eta), \qquad (3.23)$$

where  $I(\eta) = \int_{0}^{1-\eta} u^{2\alpha-1}(1-a)^{-\alpha} du$ . Finally, it follows from (3.13), (3.14) and (3.23) that

$$\sum_{2} \sim \frac{2\alpha^{3}n^{\alpha}}{\Gamma^{2}(1-\alpha)\Gamma(2\alpha+1)l(n)}I(\eta).$$
(3.24)

We now turn to estimating  $\sum_3$ . Evidently,

$$\sum_{3} < \max_{(1-\eta)n < k \le n} u_k^{(2)} \sum_{(1-\eta)n < k \le n} (n-k) p_{n-k}$$

By Lemma 2.4

$$\sum_{(1-\eta)n < k \le n} (n-k) p_{n-k} \sim \sum_{1}^{[\eta n]} \frac{l(j)}{j^{\alpha}} \sim \frac{l(n)}{1-\alpha} (\eta n)^{1-\alpha}$$

On the other hand, in view of (2.24)

$$\max_{(1-\eta)n< k\leq n} u_k^{(2)} < \frac{2^{1-\alpha}n^{\alpha}}{\Gamma(\alpha+1)} \max_{(1-\eta)n< k\leq n} \frac{\theta_k}{L(k)} \max_{(1-\eta)n/2\leq j\leq n} u_j.$$

As a result we obtain that

$$\sum_{3} = n\psi(n)(2\eta)^{1-\alpha} \max_{\delta n \le j \le n} u_j, \qquad (3.25)$$

where

$$\psi(n) = \frac{\alpha b_n}{\Gamma(\alpha+1)\Gamma(1-\alpha)(1-\alpha)}, \quad 0 < \limsup_{n \to \infty} b_n \le 1, \quad \delta = \frac{1-\eta}{2}.$$

Notice that

$$\frac{\alpha}{\Gamma(\alpha+1)\Gamma(1-\alpha)} = \frac{1}{\Gamma(\alpha)\Gamma(1-\alpha)} = \frac{\sin \pi \alpha}{\pi}$$

(see [14], formula 8.334, 3). Consequently,

$$\psi(n) = \frac{\sin \pi \alpha}{(1 - \alpha)\pi} b_n. \tag{3.26}$$

It follows from (3.1), (3.5), (3.24) and (3.25) that

$$u_n = \varphi(n) + (2\eta)^{1-\alpha} \psi(n) \max_{\delta n \le j \le n} u_j, \qquad (3.27)$$

where

$$\varphi(n) = \frac{2\alpha^3 a_n n^{\alpha-1} I(\eta)}{\Gamma^2(1-\alpha)\Gamma(2\alpha+1)l(n)}, \ a_n \sim 1.$$

Let us fix  $0 < \varepsilon < 1/2$ . Let  $\eta$  be such that  $(2\eta)^{1-\alpha} < \varepsilon$ . Chose N so that  $\psi(n) < 1$  for n > N. Let  $n_1$  be the value of k for which  $\max_{\substack{\delta n \le k \le n}} u_k$  is attained. In particular, it may be that  $n_1 = n$ . In this case  $u_n < \varphi(n)/(1-\varepsilon)$ . If  $N < n_1 < n$ , then

$$u_{n_1} < \varphi(n_1) + \varepsilon \max_{\delta n_1 \le j \le n_1} u_j$$

and consequently

$$u_n < \varphi(n) + \varepsilon \varphi(n_1) + \varepsilon^2 \max_{\delta n_1 \le j \le n_1} u_j.$$
(3.28)

If  $\max_{\delta n_1 \leq j \leq n_1} u_j = u_{n_1}$ , then  $u_{n_1} < \varphi(n_1)/(1-\varepsilon)$ . Substituting this bound in (3.28), we have

$$u_n < \varphi(n) + \varepsilon \varphi(n_1) + \frac{\varepsilon^2}{1 - \varepsilon} \varphi(n_1).$$

If  $\max_{\delta n_1 \leq j \leq n_1} u_j$  is attained for  $N < j < n_1$ , then, similarly, the following inequality is deduced

$$u_n < \varphi(n) + \varepsilon \varphi(n_1) + \varepsilon^2 \varphi(n_2) + \frac{\varepsilon^3}{1 - \varepsilon} \max_{\delta n_2 \le j \le n_2} u_j$$

and so forth.

There exist two possibilities: either for some  $n_k > N$ 

$$\max_{\delta n_k \le j \le n_k} u_j = u_{n_k},$$

or for some  $k = k_0$  the inequality  $n_k < N$  is fulfilled. Consider the first case. First of all, notice that  $n_k \ge \delta^k n$ . Using Karamata's representation (2.4) for l(n), we obtain

$$\frac{\varphi(n_j)}{\varphi(n)} = \frac{a_n a(n)}{a_{n_j} a(n_j)} \left(\frac{n}{n_j}\right)^{1-\alpha} \exp\left\{-\int_{n_j}^{n} \frac{\varepsilon(u)}{u}\right\}.$$

Evidently,

$$\left|\int\limits_{n_j}^n \frac{\varepsilon(u)}{u} du\right| < \sup_{n_j \le u \le n} |\varepsilon(u)| \ln \frac{n}{n_j} < -j\gamma \ln \delta, \ \gamma = \sup_{u > N} |\varepsilon(u)|.$$

Consequently, there exists  $\varepsilon_0$  such that for  $\varepsilon < \varepsilon_0$ 

$$\varepsilon^{j}\varphi(n_{j}) < \varepsilon^{j}\varphi(n) \exp\left\{j\gamma \ln 2\right\} < \varepsilon^{j/2}.$$

As a result we get that for  $\varepsilon < \varepsilon_0$ 

$$u_n < \sum_{j=0}^{k-1} \varepsilon^j \varphi(n_j) + \frac{\varepsilon^k}{1-\varepsilon} \varphi(n_k) < \left(\sum_{j=0}^{k-1} \varepsilon^{j/2} + \frac{\varepsilon^{k/2}}{1-\varepsilon}\right) \varphi(n) < \frac{\varphi(n)}{1-\varepsilon^{1/2}}.$$
 (3.29)

In the second case the recursion stops for  $k = k_0 = \min\{k : n_k < N\}$ . As a result we arrive at the bound

$$u_n < \frac{\varphi(n)}{1 - \varepsilon^{1/2}} + \frac{\varepsilon^{k_0 - 1}}{1 - \varepsilon} \max_{k \ge 0} u_k.$$
(3.30)

Since  $n_k \geq \delta^k n$ ,  $k_0 \geq \log_{\delta} \frac{N}{n}$ . It implies that  $\varepsilon^{k_0} \leq \exp\{-2^{-1}\ln\varepsilon\log_{\delta}n\}$  for  $n > N^2$ . Consequently, for sufficiently small  $\varepsilon$ 

$$\varepsilon^{k_0} = o(n^{-2}) = o(\varphi(n)). \tag{3.31}$$

It follows from (3.30) and (3.31) that  $u_n < 2\varphi(n)$  for  $n > N^2$  if  $\varepsilon$  sufficiently small. Returning to (3.27) we conclude that for sufficiently large n

$$0 < l(n)n^{1-\alpha}u_n - a_n c_1(\alpha)I(\eta) < 2\varepsilon n^{1-\alpha}l(n) \max_{\delta n \le k \le n} \varphi(k),$$

where  $c_1(\alpha) = 2\alpha^3/\Gamma^2(1-\alpha)\Gamma(2\alpha+1)$ . It is easily seen that

$$\limsup_{n \to \infty} n^{1-\alpha} l(n) \max_{\delta n \le k \le n} \varphi(k) \le \delta^{\alpha-1} c_1(\alpha) I(\eta).$$

It follows from two latter relations that

$$\lim_{n \to \infty} l(n) n^{1-\alpha} u_n = c_1(\alpha) I(0).$$
(3.32)

It remains to calculate  $c_1(\alpha)I(0)$ . Obviously,

$$I(0) = B(2\alpha, 1 - \alpha) = \frac{\Gamma(2\alpha)\Gamma(1 - \alpha)}{\Gamma(1 + \alpha)}$$

Consequently,

$$c_1(\alpha)I(0) = \frac{2\alpha^3\Gamma(2\alpha)}{\Gamma(1-\alpha)\Gamma(2\alpha+1)\Gamma(1+\alpha)} = \frac{\alpha}{\Gamma(1-\alpha)\Gamma(\alpha)} = \frac{\alpha\sin\pi\alpha}{\pi}.$$
 (3.33)

It follows from (3.32) and (3.33) that

$$\lim_{n \to \infty} l(n) n^{1-\alpha} u_n = \frac{\alpha \sin \pi \alpha}{\pi}$$

On the other hand, by (2.12)

$$\frac{\mathbf{P}(X=n)}{\mathbf{P}^2(X\geq n)}\sim \frac{\alpha^2}{l(n)n^{1-\alpha}}.$$

Hence,

$$\frac{\sin \pi \alpha}{\pi \alpha} \frac{\mathbf{P}(X=n)}{\mathbf{P}^2(X \ge n)} \sim \frac{\alpha \sin \pi \alpha}{\pi l(n)n^{1-\alpha}} \sim u_n,$$

which was to be proved.

# 4. The proof of Theorem 1.2

According to definition

$$h_n = C_n(-\ln(1 - f(z))).$$

Hence,

$$nh_n = C_n \left(\frac{f'(z)}{1 - f(z)}\right).$$

Consequently,

$$h_n = \frac{1}{n} \sum_{k=0}^n (k+1) p_{k+1} u_{n-k}.$$
(4.1)

Applying Theorem 1.1, we have

$$\sum_{\varepsilon n \le k \le (1-\varepsilon)n} (k+1)p_{k+1}u_{n-k} \sim \frac{\alpha \sin \pi \alpha}{\pi} \sum_{\varepsilon n \le k \le (1-\varepsilon)n} (k+1)^{-\alpha} (n-k)^{\alpha-1}$$

$$\sim \frac{\alpha \sin \pi \alpha}{\pi} \int_{\varepsilon}^{1-\varepsilon} u^{-\alpha} (1-u)^{\alpha-1} du \equiv \frac{\alpha \sin \pi \alpha}{\pi} I(\varepsilon).$$
(4.2)

On the other hand, applying Lemmas 2.4 and 2.7, we have

$$\limsup_{n \to \infty} \sum_{0 \le k < \varepsilon n} (k+1) p_{k+1} u_{n-k} < \frac{\alpha}{\pi (1-\alpha)} \left(\frac{\varepsilon}{1-\varepsilon}\right)^{1-\alpha}$$
(4.3)

and

$$\limsup_{n \to \infty} \sum_{(1-\varepsilon)n < k \le n} (k+1) p_{k+1} u_{n-k} < \frac{1}{\pi} \left(\frac{\varepsilon}{1-\varepsilon}\right)^{\alpha}.$$
(4.4)

It follows from (4.2)-(4.4) that

$$\lim_{n \to \infty} \sum_{k=0}^{n} (k+1) p_{k+1} u_{n-k} = \alpha \frac{\sin \pi \alpha}{\pi} I(0).$$
(4.5)

Obviously,

$$I(0) = B(\alpha, 1 - \alpha) = \Gamma(\alpha)\Gamma(1 - \alpha) = \frac{\pi}{\sin \pi \alpha}.$$
(4.6)

Combining (4.1), (4.5), (4.6), we obtain that

$$h_n \sim \frac{\alpha}{n},$$

which was to be proved.

Acknowledgments. I thank the referee for helpful remarks.

#### References

- Garsia A., Lamperti J. A discrete renewal theorem with infinite mean. Commentarii Mathematici Helvetici. 1963, 37, 221–234.
- De Bruijn N.G., Erdos P. On a recursion formula and some Tauberian theorems. J.Res.Nat.Bur. Stand., 1953, 50, 161–164.
- [3] Williamson J.A. Random walks and Riesz kernels. Pac. J. Math., 1968, 25, No 2, 393–415.
- [4] Rvacheva E.L. On domains of attraction of multi dimensional distributions. Selected Translations in Mathematical Statistics and Probability, 1962, 183-203. L'vov Gos. Univ. Uch. Zap. 29, Ser. Meh.– Mat., 1954, 6, No 29, p.5–44.
- [5] Nagaev S.V. Renewal theorem in the absence of power moments. Teor. Veroyatn. i Primen., 2011, 56, No 1, 188–197.
- [6] Doney R.A. One-sided local large deviation and renewal theorems in the case of infinite mean. Probab. Theory Related Fields. –Springer-Verlag, 1997, 107, 1997, 451-465.
- [7] Nagaev S. V. Large deviations of sums of independent random variables. Ann. Prob., 1979, 7, No 5, 745-789.
- [8] Seneta E. Regularly varying functions. Lecture Notes in Mathematics 508, Springer Verlag Berlin Heidelberg - New York, 1976.
- [9] Greenwood P., Omey I., J.L. Teugels J. L. Harmonic renewal measures.- Z. Warsch. Verw. Gebiete, 1982, 59, 391-409.
- [10] Grubel R.J. On harmonic renewal measures. Probab. Theory Rel. Fields, 1986, 71, 393–403.
- [11] Grubel R.J. Harmonic renewal sequences and the first positive sum. J. London Math. Soc. 1988, 38, No 2, 179–192.
- [12] Stam A.J. Some theorems on harmonic renewal measures. Stochastic processes and their Appl., 1991, 39, 277–285
- [13] Feller W. An Introduction to Probability Theory and Applications. John Wiley and Sons, New York, London, Sydney, Toronto, 1971, 2, 752p.
- [14] Gradshtein I., Ryzhik I.M. Tables of integrals, sums and products.- Fiz.Mat.Giz., Moscow, 1962, 1100 p.